

ECE 313: Exam II Conflict

Tuesday, April 9, 2024

7:00 p.m. — 8:15 p.m.

1. (a) Let X denote the number of applications with critical bugs out of the 3 samples. Let H_0 denote the hypothesis that the company's claim is true, and H_1 denote the hypothesis that the software testers' hypothesis is true. If H_0 is true, $X \sim \text{Bi}(3, \frac{1}{5})$; otherwise, if H_1 is true, $X \sim \text{Bi}(3, \frac{2}{5})$. The likelihood matrix is then: The ML decision rule is thus

	$X = 0$	$X = 1$	$X = 2$	$X = 3$
H_0	64/125	48/125	12/125	1/125
H_1	27/125	54/125	36/125	8/125

to declare H_0 if $X = 0$ and declare H_1 if $X = 1, 2, 3$.

- (b) According to your definition of H_0 and H_1 , your $p_{false\ alarm}$ may equal the value of p_{miss} in this solution, and vice versa. As long as your definitions are consistent, however, your p_e should be the same.

$$p_{false\ alarm} = P(\text{declare } H_1 | H_0) \quad (1)$$

$$= P(X = 1, 2, 3 | H_0) \quad (2)$$

$$= \frac{61}{125} \quad (3)$$

$$p_{miss} = P(\text{declare } H_0 | H_1) \quad (4)$$

$$= P(X = 0 | H_1) \quad (5)$$

$$= \frac{27}{125} \quad (6)$$

$$p_e = p_{false\ alarm} \cdot \frac{1}{3} + p_{miss} \cdot \frac{2}{3} \quad (7)$$

$$= \frac{61}{375} + \frac{18}{125} \quad (8)$$

$$= \frac{23}{75} \quad (9)$$

(c) The MAP joint probability matrix is then:

	$X = 0$	$X = 1$	$X = 2$	$X = 3$
H_0	$64/375$	$48/375$	$12/375$	$1/375$
H_1	$54/375$	$108/375$	$72/375$	$16/375$

Therefore H_0 is chosen when $X = 0$ and H_1 is chosen $X = 1, 2, 3$

(d) We notice that

$$P_{false} = \mathbf{Pr}\{H_1 \text{ claimed}|H_0\} = \mathbf{Pr}\{X \in \{1, 2, 3\}|H_0\} = \frac{48 + 12 + 1}{125} = \frac{61}{125}$$

and

$$P_{miss} = \mathbf{Pr}\{H_0 \text{ claimed}|H_1\} = \mathbf{Pr}\{X \in \{0\}|H_1\} = \frac{27}{125}$$

Thus

$$P_e = \pi_0 \cdot P_{false} + \pi_1 \cdot P_{miss} = \frac{115}{375} = \frac{23}{75}$$

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It is reasonable that this equals the p_e in (b) since the choices of hypotheses for different decision rules are the same.

2. (a)

$$\begin{aligned} P(N_{60} = 10) &= \frac{e^{-60\lambda}(60\lambda)^{10}}{10!} \\ &= \frac{180^{10}}{10!}e^{-180} \end{aligned}$$

(b)

$$\begin{aligned} &P(N_{10} = 0 \cap N_{30} - N_{10} = 0 \cap N_{40} - N_{30} = 0) \\ &= P(N_{10} = 0)P(N_{30} - N_{10} = 0)P(N_{40} - N_{30} = 0) \\ &= \frac{e^{-10\lambda}(10\lambda)^0}{0!} \frac{e^{-20\lambda}(20\lambda)^0}{0!} \frac{e^{-10\lambda}(10\lambda)^0}{0!} \\ &= e^{-120} \end{aligned}$$

or noticing that $\{N_{10} = 0 \cap N_{30} - N_{10} = 0 \cap N_{40} - N_{30} = 0\} \iff \{N_{40} = 0\}$,

$$\begin{aligned} &P(N_{10} = 0 \cap N_{30} - N_{10} = 0 \cap N_{40} - N_{30} = 0) \\ &= P(N_{40} = 0) \\ &= \frac{e^{-40\lambda}(40\lambda)^0}{0!} \\ &= e^{-120} \end{aligned}$$

(c) Let A denote the event that the Geiger counter recorded one arrival in each of the three time intervals $[0, 3]$, $[1, 4]$, and $[2, 5]$ seconds. There are three ways to meet such requirements:

	$[0, 1]$	$[1, 2]$	$[2, 3]$	$(3, 4]$	$(4, 5]$
Case #1	1	0	0	1	0
Case #2	0	1	0	0	1
Case #3	0	0	1	0	0

$$\begin{aligned} P(A) &= 2 \times \left(\frac{e^{-\lambda}\lambda^1}{1!}\right)^2 \left(\frac{e^{-\lambda}\lambda^0}{0!}\right)^3 + \left(\frac{e^{-\lambda}\lambda^1}{1!}\right)^1 \left(\frac{e^{-\lambda}\lambda^0}{0!}\right)^4 \\ &= 21e^{-15} \end{aligned}$$

3. (a) We have

$$(\text{number of times a number } > 4 \text{ shows}) = 100 - X.$$

Therefore

$$A = \{X \leq 162 - X + 4\} = \{X \leq 83\}.$$

Since a number ≤ 4 (success) occurs with probability $2/3$ on every die roll, $X \sim \text{Binom}(162, 2/3)$.

(b) Since $X \sim \text{Binom}(162, 2/3)$, therefore

$$P(A) = P\{X \leq 83\} = \sum_{k=0}^{83} \binom{162}{k} \left(\frac{2}{3}\right)^k \left(\frac{1}{3}\right)^{162-k} = \left(\frac{1}{3}\right)^{162} \sum_{k=0}^{83} \binom{162}{k} 2^k.$$

(c) Note that $E[X] = 108$, $\text{Var}(X) = (162)p(1-p) = 36$, and $\sigma = \sqrt{\text{Var}(X)} = 6$. Therefore

$$\begin{aligned} P(A) &= P\{X \leq 83\} \approx P\{\tilde{X} \leq 83.5\} \\ &= P\left\{\frac{\tilde{X} - 108}{6} \leq \frac{83.5 - 108}{6}\right\} \\ &= \Phi\left(\frac{83.5 - 108}{6}\right) = \Phi\left(-\frac{24.5}{6}\right) = \Phi(-4.08) = Q(4.08). \end{aligned}$$

where the approximation comes from Gaussian approximation with continuity correction.

(d) We have

$$\begin{aligned} P(B) &= P\{X = 55\} = P\{54.5 < X < 55.5\} \\ &\approx P\{54.5 < \tilde{X} < 55.5\} \\ &= P\left\{\frac{54.5 - 108}{6} < \frac{\tilde{X} - 108}{6} < \frac{55.5 - 108}{6}\right\} \\ &= Q\left(\frac{54.5 - 108}{6}\right) - Q\left(\frac{55.5 - 108}{6}\right) = Q(-8.91) - Q(-8.75) \\ &= \Phi(8.91) - \Phi(8.75). \end{aligned}$$

4. (a) $f_X(x)$ must be non-negative for all real x . Because c is positive, $x(1-x)$ must be nonnegative for $x \in (a, b)$.

$$x(1-x) \geq 0 \quad (10)$$

$$x \geq x^2 \quad (11)$$

This is true when $x \in [0, 1]$. Therefore, for the *pdf* to be nonnegative,

$$0 \leq a < b \leq 1$$

(b)

$$\int_a^b cx(1-x)dx = \int_a^b cx - cx^2dx \quad (12)$$

$$= c\left(\frac{x^2}{2} - \frac{x^3}{3}\right)\Big|_a^b \quad (13)$$

$$= c\left(b^2\left(\frac{1}{2} - \frac{b}{3}\right) - a^2\left(\frac{1}{2} - \frac{a}{3}\right)\right) \quad (14)$$

We know this PDF must integrate to 1 over its support, so setting this equal to 1 gives:

$$c = \left(b^2\left(\frac{1}{2} - \frac{b}{3}\right) - a^2\left(\frac{1}{2} - \frac{a}{3}\right)\right)^{-1}$$

(c)

$$\int_{\frac{1}{2}}^{\infty} f_X(x)dx = \int_{\frac{1}{2}}^b cx - cx^2dx \quad (15)$$

$$= c\left(\frac{x^2}{2} - \frac{x^3}{3}\right)\Big|_{\frac{1}{2}}^b \quad (16)$$

$$= c\left(b^2\left(\frac{1}{2} - \frac{b}{3}\right) - \frac{1}{4}\left(\frac{1}{2} - \frac{1}{6}\right)\right) \quad (17)$$

$$= c\left(b^2\left(\frac{1}{2} - \frac{b}{3}\right) - \frac{1}{12}\right) \quad (18)$$

$$(19)$$

Substituting the value of c found before gives

$$P\left(X > \frac{1}{2}\right) = \frac{b^2\left(\frac{1}{2} - \frac{b}{3}\right) - \frac{1}{12}}{b^2\left(\frac{1}{2} - \frac{b}{3}\right) - a^2\left(\frac{1}{2} - \frac{a}{3}\right)}$$